



**Streamlining the Applied Mathematics Studies
at Faculty of Science of Palacký University in Olomouc
CZ.1.07/2.2.00/15.0243**



EUROPEAN UNION



MINISTRY OF EDUCATION,
YOUTH AND SPORTS



**OP Education
for Competitiveness**

INVESTMENTS
IN EDUCATION
DEVELOPMENT

International Conference Olomoucian Days of Applied Mathematics

ODAM 2011

Department of Mathematical analysis
and Applications of Mathematics
Faculty of Science
Palacký University Olomouc

ODAM'2011

THE TYPE A UNCERTAINTY

Kubáček, L., Tesaříková, E.

FORMULATION OF THE PROBLEM

**ESTIMATION OF THE TYPE A UNCERTAINTY
IN A LINEAR MODEL**

**INSENSITIVITY REGION FOR A DISPERSION
OF THE ESTIMATOR OF LINEAR FUNCTIONS**

**LINEARIZATION REGION FOR THE BIAS OF THE
ESTIMATOR**

NUMERICAL EXAMPLE

FORMULATION OF THE PROBLEM

The following model is considered

$$\begin{pmatrix} \widehat{\Theta} \\ \mathbf{Y} \end{pmatrix} \sim N_{l+n} \left[\begin{pmatrix} \mathbf{I} & \mathbf{0} \\ \mathbf{D} & \mathbf{X} \end{pmatrix} \begin{pmatrix} \Theta \\ \beta \end{pmatrix}, \begin{pmatrix} \mathbf{W} & \mathbf{0} \\ \mathbf{0} & \Sigma \end{pmatrix} \right]$$

- Θ ... the parameter of the 1st stage,
- β ... the parameter of the 2nd stage,
- \mathbf{W} ... the type B uncertainty.

The BLUE of the parameter β is

$$\widehat{\beta}(\mathbf{Y}, \widehat{\Theta}) = [\mathbf{X}'(\Sigma + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1}\mathbf{X}'(\Sigma + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}(\mathbf{Y} - \mathbf{D}\widehat{\Theta})$$

and its covariance matrix

$$\begin{aligned} \text{Var}[\widehat{\beta}(\mathbf{Y}, \widehat{\Theta})] &= [\mathbf{X}'(\Sigma + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1} \\ &= (\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1} + (\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}\mathbf{X}'\Sigma^{-1}\mathbf{D}\mathbf{W}\mathbf{D}'\Sigma^{-1}\mathbf{X}(\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}. \end{aligned}$$

The type A uncertainty = $(\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}$.

The type B uncertainty = $(\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}\mathbf{X}'\Sigma^{-1}\mathbf{D}\mathbf{W}\mathbf{D}'\Sigma^{-1}\mathbf{X}(\mathbf{X}'\Sigma^{-1}\mathbf{X})^{-1}$.

If $\Sigma = \sigma^2\mathbf{V}$ the problem is to estimate σ^2 on the basis \mathbf{Y} resp. $\mathbf{Y} - \mathbf{D}\widehat{\Theta}$.

Assumption: the model is regular and \mathbf{W} is known matrix.

ESTIMATION OF THE TYPE A UNCERTAINTY IN LINEAR MODELS

The estimator based on \mathbf{Y} :

$$\hat{\sigma}_1^2 = \mathbf{Y}'(\mathbf{M}_{(D,X)}\mathbf{V}\mathbf{M}_{(D,X)})^+\mathbf{Y}/[n - r(\mathbf{D}, \mathbf{X})]$$

$$\text{Var}(\hat{\sigma}_1^2) = \frac{2\sigma^4}{n - r(\mathbf{D}, \mathbf{X})}.$$

need not exist ! ($n = r(\mathbf{D}, \mathbf{X})$)

The estimator based on $\mathbf{Y} - \mathbf{D}\hat{\Theta}$:

$$\hat{\sigma}_2^2 = \frac{A - B}{\text{Tr}[(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}]}$$

$$\boldsymbol{\Sigma}_0 = \sigma_0^2\mathbf{V} + \mathbf{D}\mathbf{W}\mathbf{D}',$$

$$A = (\mathbf{Y} - \mathbf{D}\hat{\Theta})'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+(\mathbf{Y} - \mathbf{D}\hat{\Theta}),$$

$$B = \text{Tr}[(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{D}\mathbf{W}\mathbf{D}']$$

$$\text{Var}_{\sigma_0^2}(\hat{\sigma}_2^2) = \frac{2}{\text{Tr}[(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}]}$$

The estimator based on $\mathbf{Y} - \mathbf{D}\hat{\Theta}$:

$$\hat{\sigma}_3^2 = (\mathbf{Y} - \mathbf{D}\hat{\Theta})'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+(\lambda_1\mathbf{V} + \lambda_2\mathbf{D}\mathbf{W}\mathbf{D}')(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+(\mathbf{Y} - \mathbf{D}\hat{\Theta})$$

$$\text{Var}_{\sigma_0^2}(\hat{\sigma}_3^2) = 2(1, 0)\mathbf{S}_{V,D\mathbf{W}\mathbf{D}'}^{-1} \begin{pmatrix} 1 \\ 0 \end{pmatrix},$$

where

$$\mathbf{S}_{V,D\mathbf{W}\mathbf{D}'} \begin{pmatrix} \lambda_1 \\ \lambda_2 \end{pmatrix} = \begin{pmatrix} 1 \\ 0 \end{pmatrix}$$

and

$$\mathbf{S}_{V,D\mathbf{W}\mathbf{D}'} = \begin{pmatrix} \boxed{\mathbf{aa}}, & \boxed{\mathbf{ab}} \\ \boxed{\mathbf{ba}}, & \boxed{\mathbf{bb}} \end{pmatrix},$$

$$\boxed{\mathbf{aa}} = \text{Tr}[\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+],$$

$$\boxed{\mathbf{ab}} = \text{Tr}[\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{D}\mathbf{W}\mathbf{D}'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+] = \boxed{\mathbf{ba}},$$

$$\boxed{\mathbf{bb}} = \text{Tr}[\mathbf{D}\mathbf{W}\mathbf{D}'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{D}\mathbf{W}\mathbf{D}'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+]$$

In the case of normality it is valid that

$$\text{Var}_{\sigma_0^2}(\hat{\sigma}_2^2) \leq \text{Var}_{\sigma_0^2}(\hat{\sigma}_3^2)$$

The estimator based on $\mathbf{Y} - \mathbf{D}\hat{\Theta}$:

$$\hat{\sigma}_4^2 = \frac{(\mathbf{Y} - \mathbf{D}\hat{\Theta})'(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+(\mathbf{Y} - \mathbf{D}\hat{\Theta}) - \text{Tr}[(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{D} \mathbf{W} \mathbf{D}']}{n - r(\mathbf{X})},$$

$$\begin{aligned} \text{Var}_{\sigma_0^2}(\hat{\sigma}_4^2) &= \frac{2\sigma^4}{n - r(\mathbf{X})} + \frac{4\sigma^2 \text{Tr}[(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{D} \mathbf{W} \mathbf{D}' (\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{V}]}{[n - r(\mathbf{X})]^2} \\ &\quad + \frac{2\text{Tr}[(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{D} \mathbf{W} \mathbf{D}' (\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{D} \mathbf{W} \mathbf{D}']}{[n - r(\mathbf{X})]^2} \end{aligned}$$

INSENSITIVITY REGION FOR THE DISPERSION OF THE ESTIMATOR OF LINEAR FUNCTIONS

Let $h(\boldsymbol{\beta}) = \mathbf{h}'\boldsymbol{\beta}$, $\boldsymbol{\beta} \in R^k$. The neighbourhood $\mathcal{N}_{h'\boldsymbol{\beta}}$ of the parameter σ_0^2 with the property

$$\sigma^2 \in \mathcal{N}_{h'\boldsymbol{\beta}} \Rightarrow \sqrt{\text{Var}_{\sigma_0^2}[\mathbf{h}'\widehat{\boldsymbol{\beta}}(\sigma^2)]} \leq (1 + \varepsilon)\sqrt{\text{Var}_{\sigma_0^2}[\mathbf{h}'\widehat{\boldsymbol{\beta}}(\sigma_0^2)]}$$

is called the insensitivity region.

In our case

$$\begin{aligned} \mathcal{N}_{h'\boldsymbol{\beta}} &= \left\{ \sigma^2 : |\sigma^2 - \sigma_0^2| \leq \right. \\ &\leq \left. \sqrt{\frac{2\varepsilon \mathbf{h}'(\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{X})^{-1}\mathbf{h}}{\mathbf{h}'(\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{X})^{-1}\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}\boldsymbol{\Sigma}_0^{-1}\mathbf{X}(\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{X})^{-1}\mathbf{h}}} \right\}. \end{aligned}$$

LINEARIZATION REGION FOR THE BIAS OF ESTIMATORS

Let instead the linear model

$$\mathbf{Y} \sim N \left[(\mathbf{D}, \mathbf{X}) \begin{pmatrix} \boldsymbol{\Theta} \\ \boldsymbol{\beta} \end{pmatrix}, \boldsymbol{\Sigma} \right],$$

a nonlinear model

$$\mathbf{Y} \sim N(\mathbf{f}(\boldsymbol{\Theta}, \boldsymbol{\beta}), \boldsymbol{\Sigma})$$

be considered and let $\mathbf{f}(\cdot, \cdot)$ can be expressed as

$$\begin{aligned} \mathbf{f}(\boldsymbol{\Theta}, \boldsymbol{\beta}) &= \mathbf{f}(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0) + \frac{\partial \mathbf{f}(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0)}{\partial \boldsymbol{\Theta}'} (\boldsymbol{\Theta} - \boldsymbol{\Theta}_0) + \frac{\partial \mathbf{f}(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0)}{\partial \boldsymbol{\beta}'} (\boldsymbol{\beta} - \boldsymbol{\beta}_0) \\ &\quad + \frac{1}{2} \begin{pmatrix} \vdots \\ \left(\begin{array}{c} \boldsymbol{\Theta} - \boldsymbol{\Theta}_0 \\ \boldsymbol{\beta} - \boldsymbol{\beta}_0 \end{array} \right)' \frac{\partial^2 f_i(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0)}{\partial \begin{pmatrix} \boldsymbol{\Theta} \\ \boldsymbol{\beta} \end{pmatrix} \partial (\boldsymbol{\Theta}', \boldsymbol{\beta}')} \begin{pmatrix} \boldsymbol{\Theta} - \boldsymbol{\Theta}_0 \\ \boldsymbol{\beta} - \boldsymbol{\beta}_0 \end{pmatrix} \\ \vdots \end{pmatrix} \\ &= \mathbf{f}_0 + \mathbf{D}\delta\boldsymbol{\Theta} + \mathbf{X}\delta\boldsymbol{\beta} + \frac{1}{2}\boldsymbol{\kappa}(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta}), \\ &\quad \delta\boldsymbol{\Theta} = \boldsymbol{\Theta} - \boldsymbol{\Theta}_0, \quad \delta\boldsymbol{\beta} = \boldsymbol{\beta} - \boldsymbol{\beta}_0, \end{aligned}$$

with a sufficiently high accuracy. Then

$$\begin{aligned} \mathbf{b} &= E(\delta\hat{\boldsymbol{\beta}}) - \delta\boldsymbol{\beta} \\ &= E(\hat{\boldsymbol{\beta}}) - \boldsymbol{\beta} = \frac{1}{2}[\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1}\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\boldsymbol{\kappa}(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta}). \end{aligned}$$

Let

$$C_b(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0) = \sup \left\{ \frac{\sqrt{\mathbf{b}'\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}\mathbf{b}}}{(\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix}} : \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in R^{l+k} \right\},$$

and at the same time

$$\begin{aligned} \mathbf{U} &= \begin{pmatrix} \boxed{\mathbf{aa}}, & \boxed{\mathbf{ab}} \\ \boxed{\mathbf{ba}}, & \boxed{\mathbf{bb}} \end{pmatrix}, \\ \boxed{\mathbf{aa}} &= \mathbf{W}, \\ \boxed{\mathbf{ab}} &= -\mathbf{W}\mathbf{D}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}[\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1} = \boxed{\mathbf{ba}}', \\ \boxed{\mathbf{bb}} &= [\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1}. \end{aligned}$$

In this case the linearization region for the bias of the estimator of the parameter $\boldsymbol{\beta}$ is

$$\mathcal{L}_b = \left\{ \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} : (\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \leq \frac{\varepsilon}{C_b(\boldsymbol{\Theta}_0, \boldsymbol{\beta}_0)} \right\}$$

and it is valid that

$$\begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in \mathcal{L}_b \Rightarrow \sqrt{\mathbf{b}'\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}\mathbf{b}} \leq \varepsilon.$$

Let

$$C_{1,\sigma^2}^{(int)} = \sup \left\{ \frac{\sqrt{\boldsymbol{\kappa}'(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta}) \left(\mathbf{M}_{(D,X)} \mathbf{V} \mathbf{M}_{(D,X)} \right)^+ \boldsymbol{\kappa}(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta})}}{(\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix}} : \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in R^{l+k} \right\}.$$

In the case of normality the linearization region for the bias of the estimator $\hat{\sigma}_1^2$ is

$$\mathcal{L}_{1,\sigma^2} = \left\{ \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} : (\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \leq \frac{\sqrt{8[n - r(\mathbf{D}, \mathbf{X})]}}{C_{1,\sigma^2}^{int}} \right\}$$

and it is valid that

$$\begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in \mathcal{L}_{1,\sigma^2} \Rightarrow \left| \sqrt{E(\hat{\sigma}_1^2)} - \sigma \right| \leq \varepsilon\sigma.$$

Let

$$C_{2,\sigma^2}^{(int)} = \sup \left\{ \frac{\sqrt{\boldsymbol{\kappa}'(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta})(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \boldsymbol{\kappa}(\delta\boldsymbol{\Theta}, \delta\boldsymbol{\beta})}}{\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}') \mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix}} : \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in R^{l+k} \right\}.$$

The linearization region for the bias of the estimator $\hat{\sigma}_2^2$ is

$$\begin{aligned} \mathcal{L}_{2,\sigma^2} &= \left\{ \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} : \right. \\ & \left. (\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}') \mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \leq \frac{\sigma}{C_{2,\sigma^2}^{(int)}} \sqrt{2\varepsilon \text{Tr}[\mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+]} \right\} \\ & \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in \mathcal{L}_{2,\sigma^2} \Rightarrow \left| \sqrt{E(\hat{\sigma}_2^2)} - \sigma \right| \leq \varepsilon\sigma. \end{aligned}$$

NUMERICAL EXAMPLE

Let, in the plane, four points A_1, A_2, A_3, A_4 be given by their coordinates, i.e. $A_i \begin{pmatrix} \Theta_{2i-1} \\ \Theta_{2i} \end{pmatrix}$, $i = 1, 2, 3, 4$,

$$A_1 \begin{pmatrix} 201.31m \\ 210.80m \end{pmatrix}, A_2 \begin{pmatrix} 406.73m \\ 863.45m \end{pmatrix}, A_3 \begin{pmatrix} 1050.47m \\ 216.66m \end{pmatrix}, A_4 \begin{pmatrix} 630.17m \\ 28.29m \end{pmatrix}.$$

The coordinates are estimated and their estimator is

$$\widehat{\Theta} = \begin{pmatrix} \widehat{\Theta}_1 \\ \vdots \\ \widehat{\Theta}_8 \end{pmatrix} \sim N_8(\Theta, \mathbf{W}), \quad \mathbf{W} = (0.1m)^2 \mathbf{I}_8.$$

Coordinates $\begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}$ of a point P must be estimated by measured distances d_i

$$d_i = E(Y_i) = \sqrt{(\Theta_{2i-1} - \beta_1)^2 + (\Theta_{2i} - \beta_2)^2}, \quad i = 1, \dots, 4,$$

where the approximate coordinates are

$$P \begin{pmatrix} 503.1m \\ 431.9m \end{pmatrix},$$

$$\mathbf{Y} \sim N_4 \left[\begin{pmatrix} d_1 \\ \vdots \\ d_4 \end{pmatrix}, (0.01m)^2 \mathbf{I}_4 \right].$$

The linearized models of this measurement are

$$\mathbf{Y} - \mathbf{f}_0 \sim N_4 \left[(\mathbf{D}, \mathbf{X}) \begin{pmatrix} \delta\Theta \\ \delta\beta \end{pmatrix}, (0.01m)^2 \mathbf{I}_4 \right]$$

and

$$\mathbf{Y} - \mathbf{f}_0 - \mathbf{D}\widehat{\delta\Theta} \sim N_4[\mathbf{X}\delta\beta, (0.01m)^2 \mathbf{I}_4 + \mathbf{D}(0.1m)^2 \mathbf{I}_8 \mathbf{D}'],$$

respectively.

Here

$$\begin{aligned}
\mathbf{f}_0 &= (f_{0,1}, \dots, f_{0,4})' = (374.12, 442.18, 588.17, 423.14)', \\
f_{0,i} &= \sqrt{(\Theta_{2i-1}^{(0)} - \beta_1^{(0)})^2 + (\Theta_{2i}^{(0)} - \beta_2^{(0)})^2}, \\
\mathbf{D} &= \left. \frac{\partial \mathbf{f}(\boldsymbol{\Theta}, \boldsymbol{\beta})}{\partial \boldsymbol{\Theta}'} \right|_{\boldsymbol{\Theta}=\boldsymbol{\Theta}^{(0)}, \boldsymbol{\beta}=\boldsymbol{\beta}^{(0)}}, \\
\frac{\partial f_{0,i}}{\partial \Theta_{2i-1}} &= \frac{\Theta_{2i-1}^{(0)} - \beta_1^{(0)}}{f_{0,i}}, \quad \frac{\partial f_{0,i}}{\partial \Theta_{2i}} = \frac{\Theta_{2i}^{(0)} - \beta_2^{(0)}}{f_{0,i}}, \\
\frac{\partial f_{0,i}}{\partial \beta_1} &= -\frac{\Theta_{2i-1}^{(0)} - \beta_1^{(0)}}{f_{0,i}}, \quad \frac{\partial f_{0,i}}{\partial \beta_2} = -\frac{\Theta_{2i}^{(0)} - \beta_2^{(0)}}{f_{0,i}}, \\
\mathbf{D} &= \begin{pmatrix} \frac{\partial f_{0,1}}{\partial \Theta_1}, & \frac{\partial f_{0,1}}{\partial \Theta_2}, & 0, & 0, & 0, & 0, & 0, & 0 \\ 0, & 0, & \frac{\partial f_{0,2}}{\partial \Theta_3}, & \frac{\partial f_{0,2}}{\partial \Theta_4}, & 0, & 0, & 0, & 0 \\ 0, & 0, & 0, & 0, & \frac{\partial f_{0,3}}{\partial \Theta_5}, & \frac{\partial f_{0,3}}{\partial \Theta_6}, & 0, & 0 \\ 0, & 0, & 0, & 0, & 0, & 0, & \frac{\partial f_{0,4}}{\partial \Theta_7}, & \frac{\partial f_{0,4}}{\partial \Theta_8} \end{pmatrix}, \\
\mathbf{D}\mathbf{D}' &= \mathbf{I}_4, \\
\mathbf{X} &= \begin{pmatrix} \frac{\partial f_{0,1}}{\partial \beta_1}, & \frac{\partial f_{0,1}}{\partial \beta_2} \\ \frac{\partial f_{0,2}}{\partial \beta_1}, & \frac{\partial f_{0,2}}{\partial \beta_2} \\ \frac{\partial f_{0,3}}{\partial \beta_1}, & \frac{\partial f_{0,3}}{\partial \beta_2} \\ \frac{\partial f_{0,4}}{\partial \beta_1}, & \frac{\partial f_{0,4}}{\partial \beta_2} \end{pmatrix} = \begin{pmatrix} 0.806\ 68, & 0.590\ 99 \\ 0.217\ 94, & -0.975\ 96 \\ -0.930\ 63, & 0.365\ 95 \\ -0.300\ 30, & 0.953\ 84 \end{pmatrix}, \\
\Sigma_0 &= (0.01m)^2 \mathbf{I}_4 + (0.1m)^2 \mathbf{D}\mathbf{D}' = 0.0101 \mathbf{I}_4.
\end{aligned}$$

The estimator $\widehat{\sigma}_1^2$ does not exist, since $\mathbf{M}_{(D,X)} = \mathbf{0}$.

The estimator $\widehat{\sigma}_2^2$ is

$$\begin{aligned}
\widehat{\sigma}_2^2 &= \frac{A - B}{\text{Tr}[(\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V} (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V}]}, \\
A &= (\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}})' (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V} (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ (\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}}), \\
B &= \text{Tr}[(\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V} (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{D}\mathbf{W}\mathbf{D}'],
\end{aligned}$$

and its dispersion is

$$\text{Var}_{\sigma_0^2}(\widehat{\sigma}_2^2) = \frac{2\sigma_0^4}{\text{Tr}[(\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V} (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ \mathbf{V}]} = 1.020\ 01 \times 10^{-4}.$$

The estimator $\widehat{\sigma}_3^2$ is

$$\widehat{\sigma}_3^2 = (\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}})' (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ (\lambda_1 \mathbf{V} + \lambda_2 \mathbf{D}\mathbf{W}\mathbf{D}') (\mathbf{M}_X \Sigma_0 \mathbf{M}_X)^+ (\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}}),$$

$$\text{Var}_{\sigma_0^2}(\widehat{\sigma}_3^2) = 2((1, 0)\mathbf{S}_{V,DWD'}^{-1} \begin{pmatrix} 1 \\ 0 \end{pmatrix}) = 1.030\ 05 \times 10^{-4},$$

where

$$\begin{aligned} \mathbf{S}_{V,DWD'} &= \begin{pmatrix} \boxed{\mathbf{aa}}, & \boxed{\mathbf{ab}} \\ \boxed{\mathbf{ba}}, & \boxed{\mathbf{bb}} \end{pmatrix}, \\ \boxed{\mathbf{aa}} &= \text{Tr}[\mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+], \\ \boxed{\mathbf{ab}} &= \text{Tr}[\mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{DWD}'(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+] = \boxed{\mathbf{ba}}, \\ \boxed{\mathbf{bb}} &= \text{Tr}[\mathbf{DWD}'(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ \mathbf{DWD}'(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+]. \end{aligned}$$

The estimator $\widehat{\sigma}_4^2$ is

$$\widehat{\sigma}_4^2 = \frac{(\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}})'(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+ (\mathbf{Y} - \mathbf{D}\widehat{\boldsymbol{\Theta}}) - \text{Tr}[(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+ \mathbf{DWD}']}{n - r(\mathbf{X})},$$

$$\begin{aligned} \text{Var}(\sigma_4^2) &= \frac{2\sigma^4}{n - r(\mathbf{X})} + \frac{4\sigma^2 \text{Tr}[(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+ \mathbf{DWD}'(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+ \mathbf{V}]}{[n - r(\mathbf{X})]^2} \\ &+ \frac{2\text{Tr}[\mathbf{DWD}'(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+ \mathbf{DWD}'(\mathbf{M}_X \mathbf{V} \mathbf{M}_X)^+]}{[n - r(\mathbf{X})]^2} = 1.020\ 01 \times 10^{-4}. \end{aligned}$$

The accuracy of different estimators of σ^2 is almost the same, even it is not sufficiently good. For example in the case of $\widehat{\sigma}^2$ it is approximately valid that

$$\frac{\sqrt{\text{Var}(\sqrt{\widehat{\sigma}^2})}}{\sigma} = \sqrt{\left(\frac{1}{2\sqrt{\sigma^2}}\right)^2 \text{Var}(\widehat{\sigma}^2)} = \frac{1}{2 \times 10^{-2}} \sqrt{1.020\ 01 \times 10^{-4}} = 0.505.$$

Thus the relative standard deviation is 50.5% what is rather large number. However nothing better can be expected because of the poor precision of the first stage measurement (\mathbf{W}).

Insensitivity regions Let

$$\mathbf{W}_h = \mathbf{h}'(\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{X})^{-1}\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0^{-1}\mathbf{M}_X)^+ \mathbf{V}\boldsymbol{\Sigma}_0^{-1}\mathbf{X}(\mathbf{X}'\boldsymbol{\Sigma}_0^{-1}\mathbf{X})^{-1}\mathbf{h}.$$

If $\mathbf{h} = (1, 0)'$, then $\mathbf{W}_{(1,0)} = \mathbf{0}$ and also for $\mathbf{h} = (0, 1)'$ $\mathbf{W}_{(0,1)} = \mathbf{0}$.

Thus the estimators of β_1 and β_2 are not sensitive on the small changes of the value σ^2 .

Linearization regions

$$\mathbf{U} = \begin{pmatrix} \boxed{\mathbf{aa}}, & \boxed{\mathbf{ab}} \\ \boxed{\mathbf{ba}}, & \boxed{\mathbf{bb}} \end{pmatrix},$$

$$\boxed{\mathbf{aa}} = \mathbf{W} = (0.1)^2 \mathbf{I}_8,$$

$$\begin{aligned} \boxed{\mathbf{ab}} &= -\mathbf{W}\mathbf{D}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}[\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1} \\ &= \begin{pmatrix} 0.004\ 533, & 0.002\ 734 \\ 0.003\ 321, & 0.002\ 003 \\ 0.000\ 091, & -0.000\ 893 \\ -0.000\ 409, & 0.003\ 998 \\ 0.005\ 089, & -0.000\ 664 \\ -0.002\ 001, & 0.000\ 261 \\ 0.000\ 287, & -0.001\ 177 \\ -0.000\ 911, & 0.003\ 738 \end{pmatrix}, \end{aligned}$$

$$\boxed{\mathbf{ba}} = -[\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1}\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{D}\mathbf{W} = \boxed{\mathbf{ab}}',$$

$$\boxed{\mathbf{bb}} = [\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}]^{-1} = \begin{pmatrix} 0.006\ 319, & 0.000\ 978 \\ 0.000\ 978, & 0.004\ 457 \end{pmatrix},$$

$$\begin{aligned} C_b(\boldsymbol{\Theta}^{(0)}, \boldsymbol{\beta}^{(0)}) &= \sup \left\{ \frac{\sqrt{\mathbf{b}'\mathbf{X}'(\boldsymbol{\Sigma} + \mathbf{D}\mathbf{W}\mathbf{D}')^{-1}\mathbf{X}\mathbf{b}}}{(\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix}} : \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in R^{l+k} \right\} \\ &= 0.000\ 119\ 76. \end{aligned}$$

If $\varepsilon = 0.1$, then

$$\mathcal{L}_b = \left\{ \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} : (\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \leq \frac{\varepsilon}{C_b(\boldsymbol{\Theta}^{(0)}, \boldsymbol{\beta}^{(0)})} \right\}$$

is the ellipsoid with the semiaxes equal to

$$\begin{aligned} a_1 &= 0.268m, & a_2 &= 0.292m, & a_3 &= 0.346m, & a_4 &= 0.346m, & a_5 &= 0.346m, \\ a_6 &= 0.346m, & a_7 &= 0.346m, & a_8 &= 0.346m, & a_9 &= 5.487m, & a_{10} &= 6.530m. \end{aligned}$$

$$C_{2,\sigma^2} = \sup \left\{ \frac{\sqrt{\boldsymbol{\kappa}'(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\mathbf{V}(\mathbf{M}_X\boldsymbol{\Sigma}_0\mathbf{M}_X)^+\boldsymbol{\kappa}}}{(\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}')\mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix}} : \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \in R^{l+k} \right\}$$

$$= -0.003\ 776.$$

If $\varepsilon = 0.1$, then

$$\begin{aligned} \mathcal{L}_{2,\sigma^2} &= \left\{ \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} : (\delta\boldsymbol{\Theta}', \delta\boldsymbol{\beta}') \mathbf{U}^{-1} \begin{pmatrix} \delta\boldsymbol{\Theta} \\ \delta\boldsymbol{\beta} \end{pmatrix} \right. \\ &\leq \left. \frac{\sigma}{C_{2,\sigma^2}} \sqrt{2\varepsilon \text{Tr}[\mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+ + \mathbf{V}(\mathbf{M}_X \boldsymbol{\Sigma}_0 \mathbf{M}_X)^+]} \right\} \end{aligned}$$

is the ellipsoid with the semiaxes

$$\begin{aligned} a_1 &= 1.504m, & a_2 &= 1.642, & a_3 &= 1.943m & a_4 &= 1.943m, & a_5 &= 1.943m, \\ a_6 &= 1.943m, & a_7 &= 1.943m, & a_8 &= 1.943m, & a_9 &= 30.813m, & a_{10} &= 36.668m. \end{aligned}$$

The linearization regions \mathcal{L}_b and \mathcal{L}_{2,σ^2} are sufficiently large with respect to requirements of geodetical practice.

Reference

- [1] Guide to the Expression of Uncertainty in Measurement. International Organization for Standardization, 1993 (Switzerland).
- [2] Korbašová, M., Marek, J.: Connecting measurements in surveying and its problem. Proceedings of INGEO and FIG Regional Central and Eastern European Conference on Engineering Surveying, Bratislava, Slovakia, November 11–13, 2004.
- [3] Kubáček, L., Kubáčková, L., Volaufová, J.: Statistical Models with Linear Structure. Veda, Bratislava 1995.
- [4] Kubáček, L., Kubáčková, L.: Statistics and Metrology (in Czech). Palacký University Olomouc, Olomouc 2000.
- [5] Kubáček, L., Marek, J.: Partial optimum estimator in two stage regression model with constraints and a problem of equivalence. *Math. Slovaca* 55, 2005, 477–494.
- [6] Marek, J.: Estimation in connecting measurements. *Acta Universitatis Palackianae, Fac. rer. nat., Mathematica* 42, 2003, 69–86.
- [7] Marek, J.: Estimation in connecting measurements with constraints of type II. *Acta Universitatis Palackianae, Fac. rer. nat., Mathematica* 43, 2004, 119–131.
- [8] Marek, J., Fišerová, E.: Statistical analysis of geodetical measurements. In: *ROBUST 2004, sborník 13. letní školy, Třešť, 7.–11. 6. 2004, JČMF Praha, 2004, 253–260.*
- [9] Rao, C. R., Mitra, S. K.: *Generalized Inverse of Matrices and Its Applications.* John Wiley & Sons, New York–London–Sydney–Toronto, 1971.